

# CONDITIONING

## CONDITIONING

### Conditioning One Random Variable on Another

**Conditional PDF** of  $X$  given that  $Y = y$ , is defined by

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$$

$y$  as a fixed number and  $f_{X|Y}(x|y)$  is as a function of the single variable  $x$ .

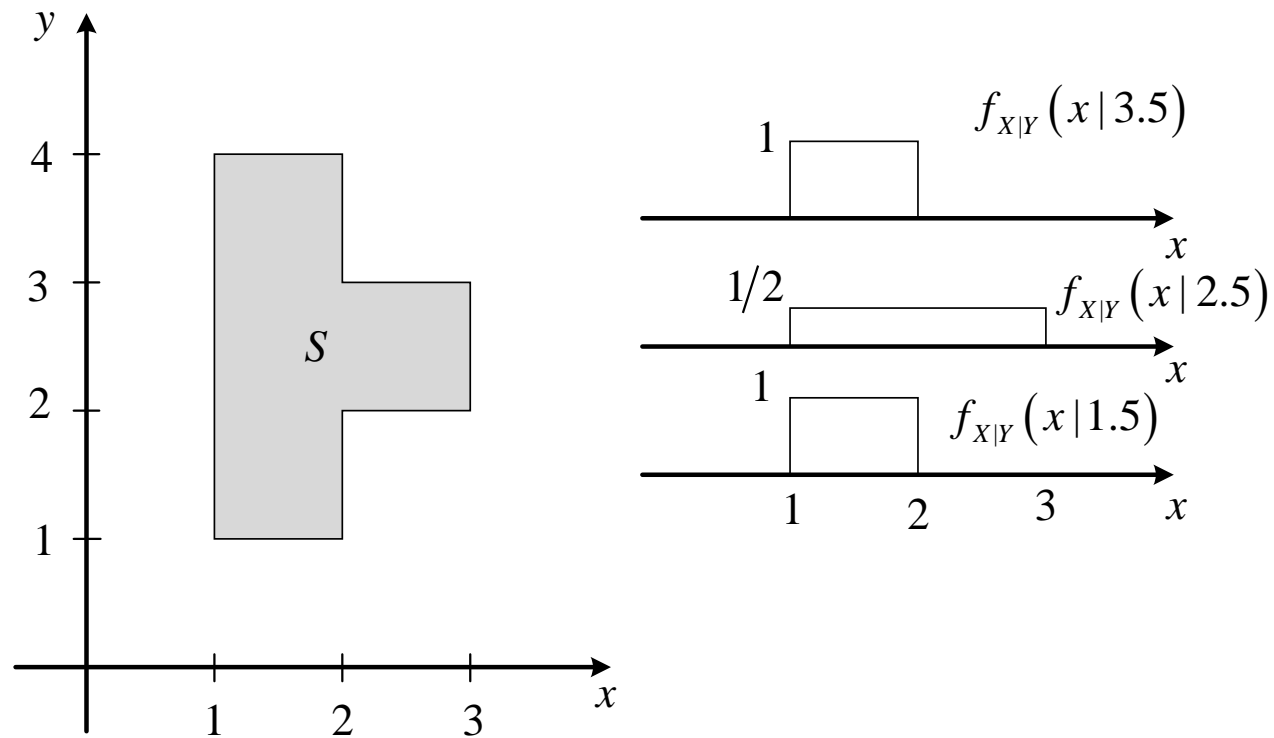
the conditional PDF  $f_{X|Y}(x|y)$  has the same shape as the joint PDF  $f_{X,Y}(x,y)$

Normalization

$$\int_{-\infty}^{\infty} f_{X|Y}(x|y) dx = 1$$

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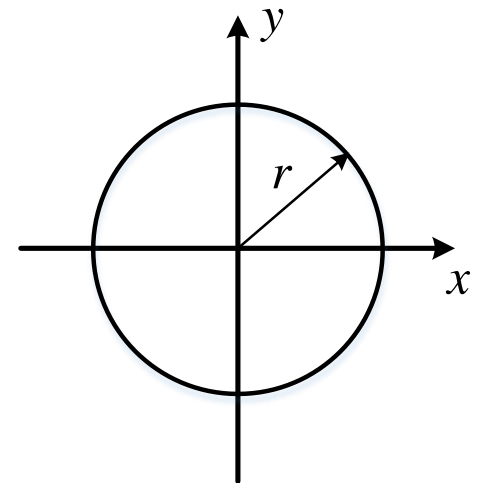


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**Example. 3.15 (textbook) Circular Uniform PDF.**

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{\text{area of the circle}}, & \text{if } (x,y) \text{ in the circle,} \\ 0, & \text{otherwise} \end{cases}$$
$$= \begin{cases} \frac{1}{\pi r^2}, & \text{if } x^2 + y^2 \leq r^2, \\ 0, & \text{otherwise} \end{cases}$$



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#### Example 3.15. (Continued)

For  $|y| > r$ ,  $f_Y(y) = 0$ . For  $|y| \leq r$ , it is given by

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx = \frac{1}{\pi r^2} \int_{x^2 + y^2 \leq r^2} dx = \frac{1}{\pi r^2} \int_{-\sqrt{r^2 - y^2}}^{\sqrt{r^2 - y^2}} dx = \frac{2}{\pi r^2} \sqrt{r^2 - y^2}, \quad |y| \leq r$$

Note that the marginal PDF  $f_Y$  is not uniform.

The conditional PDF is,

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x, y)}{f_Y(y)} = \frac{1/\pi r^2}{\frac{2}{\pi r^2} \sqrt{r^2 - y^2}} = \frac{1}{2\sqrt{r^2 - y^2}}, \quad \text{if } x^2 + y^2 \leq r^2$$

Thus, for a fixed value of  $y$ ,  $f_{X|Y}$  is uniform.

# CONDITIONING

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For more than two random variables:

$$f_{X,Y|Z}(x, y | z) = \frac{f_{X,Y,Z}(x, y, z)}{f_Z(z)}, \text{ if } f_Z(z) > 0$$

$$f_{X|Y,Z}(x | y, z) = \frac{f_{X,Y,Z}(x, y, z)}{f_{Y,Z}(y, z)}, \text{ if } f_{Y,Z}(y, z) > 0$$

Analog of the multiplication rule:

$$f_{X,Y,Z}(x, y, z) = f_{X|Y,Z}(x | y, z) f_{Y|Z}(y | z) f_Z(z)$$

# CONDITIONING

## Conditional Expectation

- **Definition:**

$$E[X | A] = \int_{-\infty}^{\infty} xf_{X|A}(x)dx$$

The conditional expectation of  $X$  given  $Y = y$  is defined by

$$E[X | Y = y] = \int_{-\infty}^{\infty} xf_{X|Y}(x|y)dx$$

- **The expected value rule:** For a function  $g(X)$ , we have

$$E[g(X) | A] = \int_{-\infty}^{\infty} g(x)f_{X|A}(x)dx \text{ and}$$

$$E[g(X) | Y] = \int_{-\infty}^{\infty} g(x)f_{X|Y}(x|y)dx$$

# CONDITIONING

## Independence

Random variables  $X$  and  $Y$  are **independent** if their joint PDF can be written as the product of the marginal PDFs:

$$f_{XY}(x, y) = f_X(x) f_Y(y) \text{ for all } x, y.$$

By using the formula  $f_{XY}(x, y) = f_{X|Y}(x|y) f_Y(y)$ , independence means that

$$f_{X|Y}(x|y) = f_X(x) \text{ for all } y \text{ with } f_Y(y) > 0 \text{ and all } x.$$

Symmetrically,

$$f_{Y|X}(y|x) = f_Y(y) \text{ for all } x \text{ with } f_X(x) > 0 \text{ and all } y.$$

# CONDITIONING

## Independence

Independence implies that

$$F_{XY}(x, y) = P(X \leq x, Y \leq y) = P(X \leq x)P(Y \leq y) = F_X(x)F_Y(y)$$

If  $X$  and  $Y$  are independent, then

$$E[g(X)h(Y)] = E[g(X)]E[h(Y)]$$

where  $g$  and  $h$  are real-valued functions.

For more than two random variables:

$$f_{XYZ}(x, y, z) = f_X(x)f_Y(y)f_Z(z) \quad \text{for all } x, y, z.$$



# CONDITIONING

## Independence

### *Summary*

- $X$  and  $Y$  are **independent** if

$$f_{XY}(x, y) = f_X(x) f_Y(y)$$

- If  $X$  and  $Y$  are **independent**, then

$$E[XY] = E[X]E[Y]$$

Furthermore, for any functions  $g$  and  $h$ , the random variables  $g(X)$  and  $h(Y)$  are independent, and therefore

$$E[g(X)h(Y)] = E[g(X)]E[h(Y)]$$

- If  $X$  and  $Y$  are **independent**, then

$$\text{var}(X + Y) = \text{var}(X) + \text{var}(Y)$$

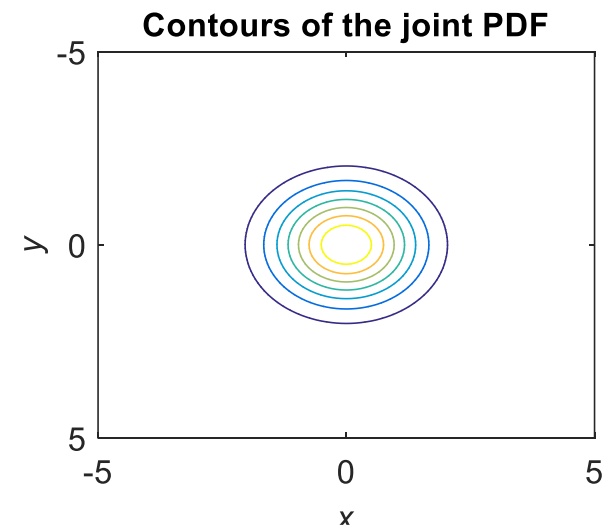
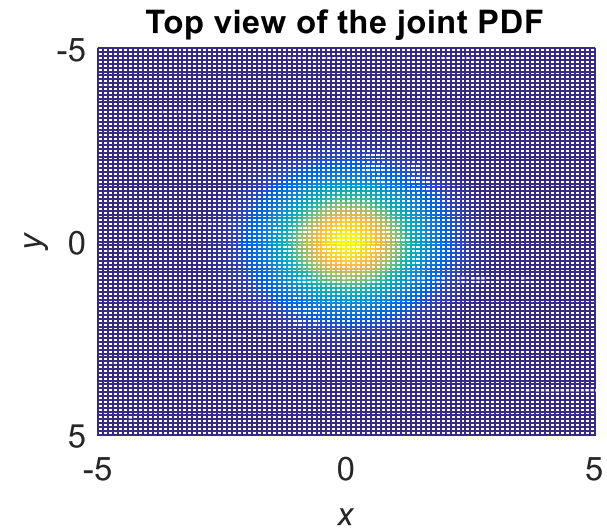
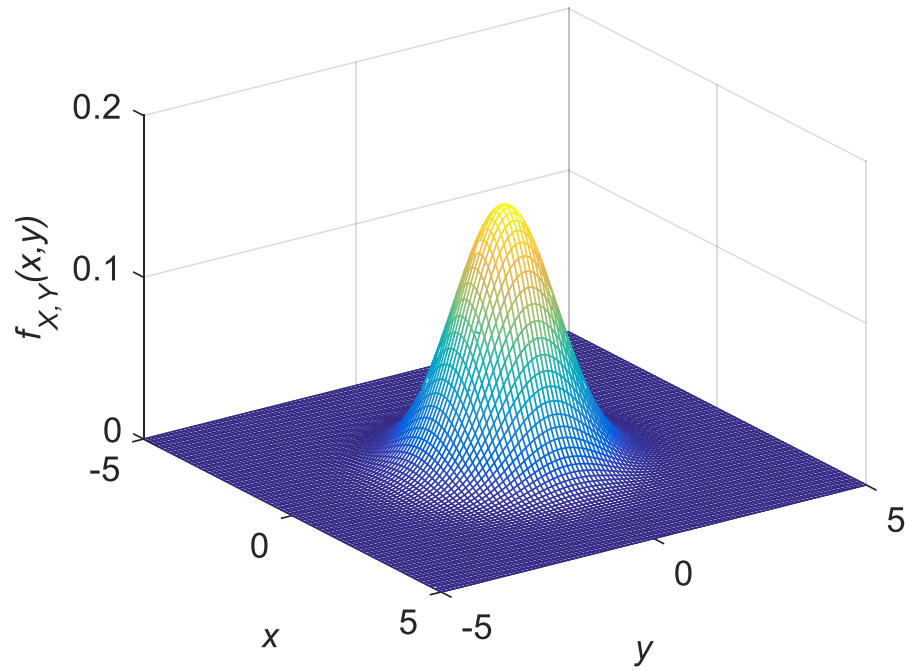
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## Independence

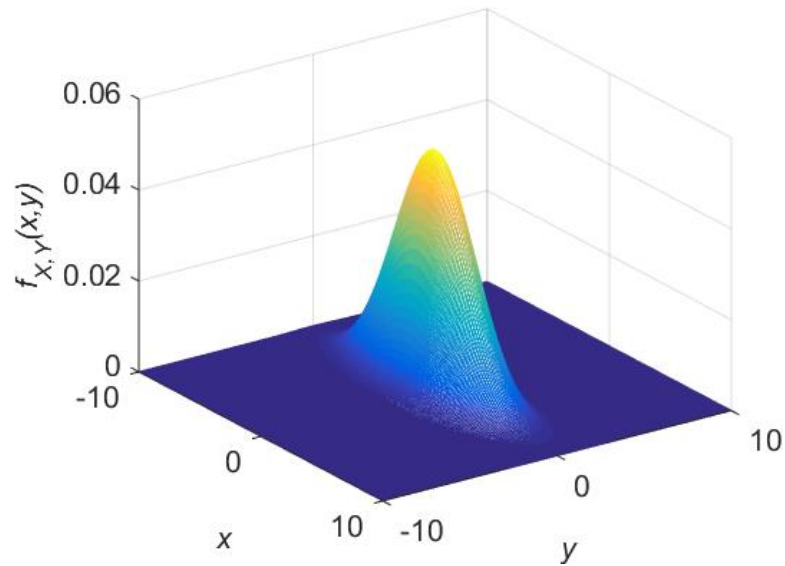
### Example 3.18 (textbook). Independent Normal Random Variables

Let  $X$  and  $Y$  independent random variables. Their joint PDF is given as

$$f_{X,Y}(x,y) = f_X(x)f_Y(y) = \frac{1}{\sqrt{2\pi\sigma_x\sigma_y}} \exp\left(-\frac{(x-\mu_x)^2}{2\sigma_x^2} - \frac{(y-\mu_y)^2}{2\sigma_y^2}\right)$$



The joint PDF of two independent Normal r.v.s in three-dimensional space ( $\sigma_X = 1, \sigma_Y = 1$ )



The joint PDF of two independent Normal r.v.s in three-dimensional space ( $\sigma_X = 3$ ,  $\sigma_Y = 1$ )

